Initial Applications of Statistical Catch-at-Age Assessment Methodology to the Gulf of Maine Winter Flounder Resource

Rebecca A. Rademeyer and Doug S. Butterworth

April 2010

Abstract

Application of SCAA to the Gulf of Maine flounder resource, though initial at this stage, suggests that with some downweighting of catch-at-age data in the likelihood, the serious retrospective problem of previous VPA assessments of this resource disappear. There are indications from the model fits considered that survey selectivity is domed (assuming commercial selectivity to be asymptotically flat at higher ages) and/or natural mortality is higher than the conventionally assumed 0.2.

Introduction

This paper presents the results of some initial applications of Statistical Catch-at-Age methodology to data for the Gulf of Maine winter flounder resource.

Data and Methodology

The catch and survey based data (including catch-at-age information) and some biological data are listed in Tables in Appendix A, from Nitschke (2011).

The details of the SCAA assessment methodology are provided in Appendix B. The Beverton-Holt stock-recruitment steepness h is fixed at 0.9 for the analyses that follow. The contribution of all catch-at-age data to the negative log-likelihood is down-weighted by a multiplicative factor W^{CAA} .

Results and Discussion

<u>Case 1</u>: Base Case with W^{CAA} =0.1, M=0.2 and commercial selectivity-at-age flat for ages 5 and above. (Figs 1-5)

Particular reasons for this choice were to not have all selectivities domed, and especially the fact that unlike the GARM3 VPA assessment (Nitschke, 2008) there is virtually no retrospective pattern (Fig. 5). Note (Fig. 1) that the spawning biomass estimates are much greater than for that GARM3 VPA. The survey selectivities are domed (Fig. 3) and fit the CAA data well, but forcing the commercial selectivity to be flat leads to systematic overestimation of the commercial plus-group numbers by the model (Fig. 4).

<u>*Case 2*</u>: Split the commercial selectivity vector estimation between 1997 and 1998 This split makes very little difference to the results; hence no plts are shown. <u>Case 3</u>: Force selectivity at age for the NEFSC surveys to be flat from age 5 and above (Fig. 6) This leads to an appreciable deterioration to the for to the data: -InL increases by 13. The primary reason for this deterioration is evident from the CAA residual plots in Fig. 6, which show a poor fit to the plus group proportions at age for the two NEFSC surveys.

<u>Case 4</u>: Fix natural mortality M = 0.4 (Fig. 7)

This leads to a 6 point improvement in the log-likelihood for the fit, with reduced residuals for the plus group for the commercial CAA data.

<u>Case 5</u>: Estimate a (constant) *M* bounded above by 0.6 (Fig. 8)

The estimated M hits the upper constraint of 0.6. There is a further improvement in the negative log-likelihood of 3 points, with the residuals for the plus group for the commercial CAA data reduced to near zero. Spawning biomass is however estimated to be lower in circumstances of an increased estimate for the pre-exploitation level.

<u>Case 6</u>: Force selectivities-at-age for all surveys to be flat above age 5 (Fig. 9) This leads to further appreciable increases in –InL, and further deterioration in the fits to the plus group proportions in the CAA for all data sets.

<u>Case 7</u>: Different weightings ($w^{survCAA}$) for the survey CAA data in the likelihood (Figs 10-12), where the reference alternative value for $w^{survCAA}$ is 0.3 (results in Table 1 are shown for this choice) in place of the 0.1 for the Base Case, but results for additional choices for $w^{survCAA}$ are shown in Fig. 11.

Results are qualitatively different for $w^{survCAA} = 0.3$, with substantial deterioration in the fits to trends in the survey abundance series (Fig. 10) and a bad retrospective pattern (Fig. 12). Fig 11 shows how as $w^{survCAA}$ is increased the fit moves closer to the VPA solution, but with a large jump between $w^{survCAA}$ values of 0.27 and 0.28 which is suggestive of a multi-modal likelihood and some conflict between the survey trend and CAA data.

Case 8: Allowance for doming in the commercial as well as the survey selectivity vectors (Fig. 13)

Unsurprisingly the negative log-likelihood improves, and the commercial plus group proportions for the CAA data are better fitted. The estimated magnitude of the spawning biomass increases markedly.

Concluding remarks

This does not pretend to be a comprehensive analysis, but some important points nevertheless seem reasonably established:

- Survey selectivity must be domed (though to a lesser extent as *M* might be set higher than 0.2).
- There is some conflict between the CAA data and the trends in the survey estimates, but if the former are given lower weight, their fit to the data does not appear visually to deteriorate substantially.
- Downweighting of the CAA data leads to higher estimated abundance, but also to the disappearance of the retrospective pattern that marks the VPA results.

References

Nitschke P. 2008. I. Gulf of Maine winter flounder. Appendix to the Report of the 3rd
Groundfish Assessment Review Meeting (GARM III): Assessment of 19 Northeast
Groundfish Stocks through 2007, Northeast Fisheries Science Center, Woods Hole,
Massachusetts,
August4-8,
4-8,
2008
http://www.nefsc.noaa.gov/publications/crd/crd0816/pdfs/garm3i.pdf

Nitschke P. 2011. Working paper (Nitschke, pers. comm)

| | 1) Ba | ise Case | (BC) | 2) Case : comme | 2: as BC ercial sel periods | but two ectivity | 3) Cas NEI selectivi | e 3: as B FSC surve ity flat fr 5 | C but eys om age | 4) Cas M=0. | se 4: as B 4 throug | 3C but shout | 5) Case | e 5: as BC estimate | C but M d | 6) Case selecti for | 6: as BC vity from all surve | but flat age 5 eys | 7) Case weight likeliho | 7: as Cas t of surve od is 0.3 of 0.1 | se 6 but ey CAA instead | 8) as BC selec | but com tivity do | mercial med |
|-------------------------------|--------------------|----------|------|--------------------|-----------------------------------|---------------------|----------------------------|--|------------------------|--------------------|------------------------|-----------------|--------------------|------------------------|--------------|---------------------------|------------------------------------|--------------------------|-------------------------------|--|-------------------------------|--------------------|----------------------|----------------|
| '-InL:overall | -123.2 | | | -123.5 | | | -110.1 | | | -129.1 | | | -132.3 | | | -101.3 | | | -156.9 | | | -133.5 | | |
| '-InL:Survey | -72.4 | | | -72.4 | | | -71.6 | | | -72.4 | | | -72.7 | | | -79.5 | | | 1.8 | | | -72.8 | | |
| '-InL:CAA | 8.0 | | | 7.4 | | | 7.2 | | | 2.3 | | | -0.1 | | | 7.8 | | | -1.6 | | | -1.9 | | |
| '-InL:CAAsurv | -42.7 | | | -42.8 | | | -29.0 | | | -42.4 | | | -42.9 | | | -14.4 | | | -142.5 | | | -42.6 | | |
| '-InL:RecRes | -17.2 | | | -17.2 | | | -17.4 | | | -17.1 | | | -17.0 | | | -15.8 | | | -14.9 | | | -17.1 | | |
| '-InL:SelSmoothing | 1.0 | | | 1.5 | | | 0.7 | | | 0.6 | | | 0.4 | | | 0.5 | | | 0.3 | | | 0.9 | | |
| | | | | | | | | | | | | | | | | | | | | | | | | |
| h | 0.90 | | | 0.90 | | | 0.90 | | | 0.90 | | | 0.90 | | | 0.90 | | | 0.90 | | | 0.90 | | |
| м | 0.20 | | | 0.20 | | | 0.20 | | | 0.40 | | | 0.60 | | | 0.20 | | | 0.20 | | | 0.20 | | |
| Theta | 0.50 | | | 0.54 | | | 0.35 | | | 0.79 | | | 0.25 | | | 0.25 | | | 0.41 | | | 0.62 | | |
| Phi | 0.12 | | | 0.11 | | | 0.19 | | | 0.08 | | | 0.34 | | | 0.30 | | | 0.30 | | | 0.15 | | |
| ρ _{sr} | - | | | 0.00 | | | 0.00 | | | 0.00 | | | 0.00 | | | 0.00 | | | 0.00 | | | 0.00 | | |
| | | | | | | | | | | | | | | | | | | | | | | | | |
| Ksp | 37.97 | | | 37.90 | | | 38.98 | | | 20.35 | | | 55.94 | | | 41.00 | | | 23.09 | | | 53.06 | | |
| B ^{sp} 2010 | 15.88 | | | 15.98 | | | 15.78 | | | 11.39 | | | 10.33 | | | 16.46 | | | 4.73 | | | 23.53 | | |
| B_{2010}^{sp}/K^{sp} | 0.42 | | | 0.42 | | | 0.40 | | | 0.56 | | | 0.18 | | | 0.40 | | | 0.20 | | | 0.44 | | |
| $B^{sp}_{2010}/B^{sp}_{1082}$ | 0.83 | | | 0.79 | | | 1.15 | | | 0.71 | | | 0.73 | | | 1.64 | | | 0.50 | | | 0.72 | | |
| - 2010/ - 1962 | | | | | | | | | | | | | | | | | | | | | | | | |
| MSYL ^{sp} | 0.17 | | | 0.17 | | | 0.17 | | | 0.15 | | | 0.15 | | | 0.17 | | | 0.17 | | | 0.14 | | |
| B ^{sp} MSY | 6.33 | | | 6.43 | | | 6.56 | | | 3.10 | | | 8.42 | | | 6.93 | | | 3.94 | | | 7.18 | | |
| MSY | 1.89 | | | 1.97 | | | 1.98 | | | 2.59 | | | 7.24 | | | 2.09 | | | 1.21 | | | 2.37 | | |
| σ _{comCAA} | 0.21 | | | 0.21 | | | 0.21 | | | 0.17 | | | 0.15 | | | 0.21 | | | 0.14 | | | 0.14 | | |
| | | | | | | | | | | | | | | | | | | | | | | | | |
| Survey | a x10 ⁶ | σ | σ | a x10 ⁶ | σ | σ | a x10 ⁶ | σ | σ | a x10 ⁶ | σ | σ | a x10 ⁶ | σ | σ | a x10 ⁶ | σ | σ | a x10 ⁶ | σ | σ | a x10 ⁶ | σ | σ |
| NEFSCspring | 0.31 | 0.55 | 0.15 | 0.31 | 0.54 | 0.15 | 0.29 | 0.57 | 0.21 | 0.20 | 0.55 | 0.16 | 0.18 | 0.54 | 0.16 | 0.27 | 0.58 | 0.22 | 0.71 | 0.63 | 0.16 | 0.22 | 0.55 | 0.15 |
| NEFSCfall | 0.45 | 0.46 | 0.16 | 0.46 | 0.46 | 0.16 | 0.42 | 0.47 | 0.23 | 0.30 | 0.46 | 0.16 | 0.24 | 0.47 | 0.16 | 0.40 | 0.46 | 0.24 | 0.94 | 0.72 | 0.15 | 0.31 | 0.46 | 0.16 |
| MADspring | 4.08 | 0.25 | 0.14 | 4.07 | 0.25 | 0.14 | 4.04 | 0.25 | 0.14 | 1.98 | 0.25 | 0.13 | 1.29 | 0.25 | 0.13 | 3.62 | 0.28 | 0.20 | 7.59 | 0.55 | 0.12 | 2.83 | 0.26 | 0.14 |
| MADfall | 4.32 | 0.17 | 0.12 | 4.33 | 0.17 | 0.12 | 4.26 | 0.17 | 0.13 | 2.26 | 0.17 | 0.13 | 1.31 | 0.17 | 0.13 | 3.89 | 0.12 | 0.19 | 8.30 | 0.58 | 0.12 | 3.01 | 0.17 | 0.13 |
| σ _{R out} | 0.27 | | | 0.27 | | | 0.24 | | | 0.29 | | | 0.30 | | | 0.29 | | | 0.33 | | | 0.29 | | |

Table 1: Results of SCAA for the Gulf of Maine winter flounder – see main text and Appendix B for specifications and definitions of some of the symbols used. Biomass units are '000t. Values input rather than estimated are shown in bold.



Fig. 1: Spawning stock biomass trajectories for the Base Case, compared to the GARM3 VPA (Nitschke, 2008).



Fig. 2: Stock-recruit relationship and estimated stock-recruit residuals for the Base Case.



Fig. 3: Commercial and survey selectivities-at-age estimated for the Base Case.



Fig. 4: The first two rows give the fit of the Base Case to the survey indices of abundance and corresponding survey standardised residuals. The third and fourth row plot the fit of the Base Case to the commercial and survey catch-at-age data. The third row compares the observed and predicted CAA as averaged over all years for which data are available, while the fourth row plots the standardised residuals, with the size (area) of the bubbles being proportional to the magnitude of the corresponding standardised residuals. For positive residuals, the bubbles are grey, whereas for negative residuals, the bubbles are white.



Fig. 5: Retrospective analysis of spawning biomass and recruitment for the Base Case.



Fig. 6: Spawning stock biomass trajectories, stock-recruit relationship, recruitment residuals and selectivities for Case 3 (NEFSC survey selectivity flat). The fits to the commercial and survey CAA are also shown.



Fig. 7: Spawning stock biomass trajectories, stock-recruit relationship, recruitment residuals and selectivities for Case 4 (M = 0.4). The fits to the commercial and survey CAA are also shown.



Fig. 8: Spawning stock biomass trajectories, stock-recruit relationship, recruitment residuals and selectivities for Case 5 (*M* estimated at 0.6). The fits to the commercial and survey CAA are also shown.



Fig. 9: Spawning stock biomass trajectories, stock-recruit relationship, recruitment residuals and selectivities for Case 6 (flat selectivities for all surveys). The fits to the commercial and survey CAA are also shown.



Fig. 10: Spawning stock biomass trajectories, stock-recruit relationship, recruitment residuals and selectivities for Case 7 (survey CAA data upweighted in the likelihood). The fits to the commercial and survey CAA and to the survey indices are also shown.



Fig. 11: Spawning stock biomass trajectories for Case 7 with different weightings (w^{CAA}) for the survey CAA data in the likelihood. The VPA results are also shown (Nitschke, 2008).



Fig. 12: Retrospective analysis of spawning biomass and recruitment for Case 7.



Fig. 13: Spawning stock biomass trajectories, stock-recruit relationship, recruitment residuals and selectivities for Case 8 (commercial selectivity domed). The fits to the commercial and survey CAA are also shown.

APPENDIX A – Data

| Voor | Total catch | Voor | Total catch |
|------|-------------|------|-------------|
| Tear | (mt) | Teal | (mt) |
| 1982 | 6178 | 1997 | 660 |
| 1983 | 3035 | 1998 | 689 |
| 1984 | 2883 | 1999 | 399 |
| 1985 | 3327 | 2000 | 587 |
| 1986 | 1692 | 2001 | 756 |
| 1987 | 2713 | 2002 | 740 |
| 1988 | 1927 | 2003 | 801 |
| 1989 | 2315 | 2004 | 687 |
| 1990 | 1511 | 2005 | 387 |
| 1991 | 1136 | 2006 | 247 |
| 1992 | 947 | 2007 | 297 |
| 1993 | 778 | 2008 | 405 |
| 1994 | 640 | 2009 | 367 |
| 1995 | 776 | 2010 | 195 |
| 1996 | 674 | | |

Table A1: Total catch (metric tons) for Gulf of Maine winter flounder (Nitschke, 2011).

| Table A2. Catch at age matrix (000s) for Gulf of Maine winter flounder (Nitschke, 2011) |
|---|
| 0 () |

| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8+ |
|------|-----|------|------|------|------|-----|-----|-----|
| 1982 | 112 | 2883 | 5267 | 3487 | 1402 | 617 | 276 | 104 |
| 1983 | 135 | 915 | 1955 | 1838 | 857 | 362 | 158 | 133 |
| 1984 | 23 | 916 | 2077 | 1901 | 856 | 348 | 312 | 225 |
| 1985 | 31 | 288 | 1598 | 2122 | 1925 | 398 | 218 | 136 |
| 1986 | 49 | 505 | 928 | 851 | 373 | 353 | 102 | 62 |
| 1987 | 53 | 486 | 2004 | 1224 | 794 | 311 | 138 | 136 |
| 1988 | 23 | 471 | 1188 | 1177 | 361 | 248 | 123 | 89 |
| 1989 | 24 | 238 | 1353 | 1478 | 777 | 213 | 51 | 38 |
| 1990 | 9 | 263 | 836 | 1008 | 504 | 172 | 49 | 29 |
| 1991 | 18 | 304 | 864 | 610 | 234 | 119 | 57 | 41 |
| 1992 | 44 | 390 | 734 | 585 | 207 | 72 | 28 | 18 |
| 1993 | 28 | 197 | 758 | 669 | 149 | 69 | 9 | 3 |
| 1994 | 18 | 81 | 503 | 623 | 152 | 44 | 16 | 7 |
| 1995 | 27 | 70 | 335 | 765 | 392 | 122 | 18 | 18 |
| 1996 | 16 | 217 | 733 | 350 | 79 | 13 | 7 | 11 |
| 1997 | 19 | 286 | 592 | 449 | 117 | 22 | 8 | 12 |
| 1998 | 20 | 64 | 264 | 474 | 333 | 115 | 41 | 12 |
| 1999 | 7 | 13 | 79 | 240 | 227 | 103 | 29 | 28 |
| 2000 | 17 | 29 | 89 | 394 | 380 | 142 | 34 | 15 |
| 2001 | 13 | 21 | 84 | 384 | 432 | 242 | 101 | 56 |
| 2002 | 4 | 31 | 167 | 383 | 408 | 187 | 65 | 34 |
| 2003 | 9 | 41 | 168 | 390 | 419 | 247 | 78 | 46 |
| 2004 | 10 | 89 | 202 | 345 | 250 | 195 | 64 | 47 |
| 2005 | 15 | 54 | 165 | 259 | 139 | 55 | 17 | 16 |
| 2006 | 7 | 14 | 104 | 160 | 89 | 27 | 14 | 12 |
| 2007 | 5 | 23 | 93 | 193 | 135 | 57 | 16 | 9 |
| 2008 | 8 | 21 | 75 | 181 | 205 | 116 | 66 | 40 |
| 2009 | 6 | 22 | 54 | 146 | 219 | 144 | 41 | 26 |
| 2010 | 6 | 10 | 20 | 70 | 120 | 84 | 40 | 16 |

| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8+ |
|------|-------|-------|-------|-------|-------|-------|-------|-------|
| 1982 | 0.084 | 0.224 | 0.375 | 0.487 | 0.595 | 0.802 | 0.943 | 2.037 |
| 1983 | 0.123 | 0.257 | 0.358 | 0.502 | 0.644 | 0.795 | 0.946 | 1.164 |
| 1984 | 0.082 | 0.264 | 0.306 | 0.401 | 0.543 | 0.708 | 0.855 | 1.115 |
| 1985 | 0.043 | 0.174 | 0.312 | 0.447 | 0.584 | 0.809 | 0.927 | 1.122 |
| 1986 | 0.050 | 0.309 | 0.410 | 0.510 | 0.664 | 0.813 | 1.005 | 1.221 |
| 1987 | 0.035 | 0.259 | 0.392 | 0.527 | 0.690 | 0.858 | 1.070 | 1.284 |
| 1988 | 0.038 | 0.396 | 0.426 | 0.487 | 0.648 | 0.754 | 1.022 | 1.204 |
| 1989 | 0.040 | 0.229 | 0.427 | 0.582 | 0.629 | 1.004 | 1.175 | 1.397 |
| 1990 | 0.034 | 0.301 | 0.421 | 0.538 | 0.625 | 0.763 | 0.979 | 1.226 |
| 1991 | 0.038 | 0.277 | 0.451 | 0.583 | 0.599 | 0.695 | 0.744 | 0.929 |
| 1992 | 0.027 | 0.227 | 0.406 | 0.533 | 0.638 | 0.788 | 1.051 | 1.465 |
| 1993 | 0.028 | 0.238 | 0.367 | 0.439 | 0.645 | 0.667 | 1.115 | 1.453 |
| 1994 | 0.028 | 0.090 | 0.369 | 0.470 | 0.610 | 0.747 | 1.068 | 1.229 |
| 1995 | 0.038 | 0.105 | 0.341 | 0.421 | 0.535 | 0.635 | 0.833 | 1.563 |
| 1996 | 0.028 | 0.321 | 0.454 | 0.541 | 0.643 | 0.722 | 0.767 | 1.321 |
| 1997 | 0.038 | 0.240 | 0.421 | 0.512 | 0.628 | 0.889 | 0.784 | 0.921 |
| 1998 | 0.029 | 0.202 | 0.392 | 0.472 | 0.615 | 0.755 | 0.910 | 1.557 |
| 1999 | 0.039 | 0.114 | 0.377 | 0.487 | 0.542 | 0.665 | 0.838 | 1.219 |
| 2000 | 0.041 | 0.146 | 0.353 | 0.473 | 0.581 | 0.698 | 0.817 | 1.03 |
| 2001 | 0.034 | 0.115 | 0.319 | 0.448 | 0.538 | 0.693 | 0.852 | 1.194 |
| 2002 | 0.050 | 0.182 | 0.415 | 0.496 | 0.593 | 0.705 | 0.882 | 1.285 |
| 2003 | 0.035 | 0.156 | 0.366 | 0.482 | 0.560 | 0.704 | 0.889 | 1.436 |
| 2004 | 0.035 | 0.207 | 0.352 | 0.494 | 0.628 | 0.763 | 0.923 | 1.269 |
| 2005 | 0.042 | 0.172 | 0.380 | 0.505 | 0.669 | 0.895 | 1.038 | 1.346 |
| 2006 | 0.048 | 0.138 | 0.404 | 0.535 | 0.715 | 0.811 | 1.032 | 1.365 |
| 2007 | 0.043 | 0.200 | 0.386 | 0.487 | 0.639 | 0.815 | 0.964 | 1.476 |
| 2008 | 0.046 | 0.153 | 0.375 | 0.474 | 0.549 | 0.671 | 0.784 | 1.097 |
| 2009 | 0.043 | 0.155 | 0.329 | 0.449 | 0.565 | 0.678 | 0.692 | 1.115 |
| 2010 | 0.031 | 0.065 | 0.314 | 0.427 | 0.507 | 0.604 | 0.717 | 0.947 |

Table A3a. Total fishery mean weights-at-age (kg) for Gulf of Maine winter flounder(Nitschke, 2011).

Table A3b. Spawning stock biomass mean weights-at-age (kg) for Gulf of Maine winter flounder (Nitschke, 2011).

| | | 2 | | | - | - | _ | - |
|------|-------|-------|-------|-------|-------|-------|-------|-------|
| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8+ |
| 1982 | 0.048 | 0.177 | 0.324 | 0.424 | 0.515 | 0.738 | 0.870 | 2.037 |
| 1983 | 0.084 | 0.147 | 0.283 | 0.434 | 0.560 | 0.688 | 0.871 | 1.164 |
| 1984 | 0.056 | 0.180 | 0.280 | 0.379 | 0.522 | 0.675 | 0.825 | 1.115 |
| 1985 | 0.016 | 0.119 | 0.287 | 0.370 | 0.484 | 0.663 | 0.810 | 1.122 |
| 1986 | 0.022 | 0.115 | 0.267 | 0.399 | 0.545 | 0.689 | 0.902 | 1.221 |
| 1987 | 0.010 | 0.114 | 0.348 | 0.465 | 0.593 | 0.755 | 0.933 | 1.284 |
| 1988 | 0.016 | 0.118 | 0.332 | 0.437 | 0.584 | 0.721 | 0.936 | 1.204 |
| 1989 | 0.015 | 0.093 | 0.411 | 0.498 | 0.554 | 0.807 | 0.941 | 1.397 |
| 1990 | 0.012 | 0.110 | 0.311 | 0.479 | 0.603 | 0.693 | 0.991 | 1.226 |
| 1991 | 0.016 | 0.097 | 0.368 | 0.495 | 0.568 | 0.659 | 0.753 | 0.929 |
| 1992 | 0.009 | 0.093 | 0.335 | 0.490 | 0.610 | 0.687 | 0.855 | 1.465 |
| 1993 | 0.016 | 0.080 | 0.289 | 0.422 | 0.586 | 0.652 | 0.937 | 1.453 |
| 1994 | 0.015 | 0.050 | 0.296 | 0.415 | 0.518 | 0.694 | 0.844 | 1.229 |
| 1995 | 0.013 | 0.054 | 0.175 | 0.394 | 0.501 | 0.622 | 0.789 | 1.563 |
| 1996 | 0.010 | 0.110 | 0.218 | 0.430 | 0.520 | 0.622 | 0.698 | 1.321 |
| 1997 | 0.017 | 0.082 | 0.368 | 0.482 | 0.583 | 0.756 | 0.752 | 0.921 |
| 1998 | 0.015 | 0.088 | 0.307 | 0.446 | 0.561 | 0.689 | 0.899 | 1.557 |
| 1999 | 0.020 | 0.058 | 0.276 | 0.437 | 0.506 | 0.640 | 0.795 | 1.219 |
| 2000 | 0.025 | 0.076 | 0.201 | 0.422 | 0.532 | 0.615 | 0.737 | 1.03 |
| 2001 | 0.015 | 0.069 | 0.216 | 0.398 | 0.505 | 0.635 | 0.771 | 1.194 |
| 2002 | 0.028 | 0.079 | 0.219 | 0.398 | 0.515 | 0.616 | 0.782 | 1.285 |
| 2003 | 0.014 | 0.088 | 0.258 | 0.447 | 0.527 | 0.646 | 0.792 | 1.436 |
| 2004 | 0.016 | 0.085 | 0.234 | 0.425 | 0.550 | 0.654 | 0.806 | 1.269 |
| 2005 | 0.023 | 0.078 | 0.281 | 0.422 | 0.575 | 0.750 | 0.890 | 1.346 |
| 2006 | 0.024 | 0.076 | 0.264 | 0.451 | 0.601 | 0.737 | 0.961 | 1.365 |
| 2007 | 0.023 | 0.098 | 0.231 | 0.444 | 0.585 | 0.763 | 0.884 | 1.476 |
| 2008 | 0.025 | 0.081 | 0.274 | 0.428 | 0.517 | 0.655 | 0.799 | 1.097 |
| 2009 | 0.035 | 0.084 | 0.224 | 0.410 | 0.518 | 0.610 | 0.681 | 1.115 |
| 2010 | 0.018 | 0.053 | 0.223 | 0.369 | 0.477 | 0.588 | 0.700 | 0.969 |

| , | | | | | | | | |
|------|-------|-------|-------|-------|-------|-------|-------|-------|
| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8+ |
| 1982 | 0.048 | 0.177 | 0.324 | 0.424 | 0.515 | 0.738 | 0.870 | 2.037 |
| 1983 | 0.084 | 0.147 | 0.283 | 0.434 | 0.560 | 0.688 | 0.871 | 1.164 |
| 1984 | 0.056 | 0.180 | 0.280 | 0.379 | 0.522 | 0.675 | 0.825 | 1.115 |
| 1985 | 0.016 | 0.119 | 0.287 | 0.370 | 0.484 | 0.663 | 0.810 | 1.122 |
| 1986 | 0.022 | 0.115 | 0.267 | 0.399 | 0.545 | 0.689 | 0.902 | 1.221 |
| 1987 | 0.010 | 0.114 | 0.348 | 0.465 | 0.593 | 0.755 | 0.933 | 1.284 |
| 1988 | 0.016 | 0.118 | 0.332 | 0.437 | 0.584 | 0.721 | 0.936 | 1.204 |
| 1989 | 0.015 | 0.093 | 0.411 | 0.498 | 0.554 | 0.807 | 0.941 | 1.397 |
| 1990 | 0.012 | 0.110 | 0.311 | 0.479 | 0.603 | 0.693 | 0.991 | 1.226 |
| 1991 | 0.016 | 0.097 | 0.368 | 0.495 | 0.568 | 0.659 | 0.753 | 0.929 |
| 1992 | 0.009 | 0.093 | 0.335 | 0.490 | 0.610 | 0.687 | 0.855 | 1.465 |
| 1993 | 0.016 | 0.080 | 0.289 | 0.422 | 0.586 | 0.652 | 0.937 | 1.453 |
| 1994 | 0.015 | 0.050 | 0.296 | 0.415 | 0.518 | 0.694 | 0.844 | 1.229 |
| 1995 | 0.013 | 0.054 | 0.175 | 0.394 | 0.501 | 0.622 | 0.789 | 1.563 |
| 1996 | 0.010 | 0.110 | 0.218 | 0.430 | 0.520 | 0.622 | 0.698 | 1.321 |
| 1997 | 0.017 | 0.082 | 0.368 | 0.482 | 0.583 | 0.756 | 0.752 | 0.921 |
| 1998 | 0.015 | 0.088 | 0.307 | 0.446 | 0.561 | 0.689 | 0.899 | 1.557 |
| 1999 | 0.020 | 0.058 | 0.276 | 0.437 | 0.506 | 0.640 | 0.795 | 1.219 |
| 2000 | 0.025 | 0.076 | 0.201 | 0.422 | 0.532 | 0.615 | 0.737 | 1.03 |
| 2001 | 0.015 | 0.069 | 0.216 | 0.398 | 0.505 | 0.635 | 0.771 | 1.194 |
| 2002 | 0.028 | 0.079 | 0.219 | 0.398 | 0.515 | 0.616 | 0.782 | 1.285 |
| 2003 | 0.014 | 0.088 | 0.258 | 0.447 | 0.527 | 0.646 | 0.792 | 1.436 |
| 2004 | 0.016 | 0.085 | 0.234 | 0.425 | 0.550 | 0.654 | 0.806 | 1.269 |
| 2005 | 0.023 | 0.078 | 0.281 | 0.422 | 0.575 | 0.750 | 0.890 | 1.346 |
| 2006 | 0.024 | 0.076 | 0.264 | 0.451 | 0.601 | 0.737 | 0.961 | 1.365 |
| 2007 | 0.023 | 0.098 | 0.231 | 0.444 | 0.585 | 0.763 | 0.884 | 1.476 |
| 2008 | 0.025 | 0.081 | 0.274 | 0.428 | 0.517 | 0.655 | 0.799 | 1.097 |
| 2009 | 0.035 | 0.084 | 0.224 | 0.410 | 0.518 | 0.610 | 0.681 | 1.115 |
| 2010 | 0.018 | 0.053 | 0.223 | 0.369 | 0.477 | 0.588 | 0.700 | 0.969 |

Table A3c. January-1 mean weights-at-age (kg) for Gulf of Maine winter flounder (Nitschke,2011).

| Table A4: Proportion mature-at-age | for Gulf of Maine winter flounder | (Nitschke, 2011) |
|------------------------------------|-----------------------------------|------------------|
|------------------------------------|-----------------------------------|------------------|

| 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8+ |
|------|------|------|------|------|------|------|------|
| 0.00 | 0.04 | 0.35 | 0.88 | 0.99 | 1.00 | 1.00 | 1.00 |

| | NEFSC spring | NEFSC fall | MADMF spring | MADMF fall |
|-------|-----------------|------------|-----------------|---------------|
| Month | 4 | 10 | 5 | 5 |
| Ages | 1-8+ | 1-8+ | 1-8+ | 1-8+ |
| 1982 | 7.67 | 4.201 | 61.61 | 108.20 |
| 1983 | 12.367 | 10.304 | 112.49 | 76.66 |
| 1984 | 5.155 | 7.732 | 68.95 | 39.54 |
| 1985 | 3.469 | 7.638 | 54.21 | 48.68 |
| 1986 | 2.342 | 2.502 | 68.98 | 44.65 |
| 1987 | 5.609 | 1.605 | 85.18 | 54.43 |
| 1988 | 6.897 | 3 | 54.04 | 38.42 |
| 1989 | 3.717 | 6.402 | 64.70 | 39.25 |
| 1990 | 5.415 | 3.527 | 82.13 | 67.66 |
| 1991 | 4.517 | 7.035 | 46.63 | 101.72 |
| 1992 | 3.932 | 10.447 | 79.00 | 87.58 |
| 1993 | 1.556 | 7.559 | 78.02 | 93.53 |
| 1994 | 3.481 | 4.87 | 72.58 | 67.79 |
| 1995 | 12.185 | 4.765 | 89.36 | 76.74 |
| 1996 | 2.736 | 10.099 | 70.49 | 77.01 |
| 1997 | 2.806 | 10.008 | 85.40 | 78.40 |
| 1998 | 2.001 | 3.218 | 77.77 | 98.45 |
| 1999 | 6.51 | 10.921 | 80.78 | 125.74 |
| 2000 | 10.383 | 12.705 | 162.19 | 99.95 |
| 2001 | 5.242 | 8.786 | 89.74 | 81.07 |
| 2002 | 12.066 | 10.691 | 91.08 | 65.81 |
| 2003 | 7.839 | 10.182 | 83.69 | 90.48 |
| 2004 | 3.879 | 2.763 | 79.12 | 107.59 |
| 2005 | 6.92 | 8.807 | 94.04 | 78.59 |
| 2006 | 4.173 | 7.117 | 85.55 | 86.99 |
| 2007 | 2.5 | 6.378 | 53.58 | 76.67 |
| 2008 | 11.543 | 13.319 | 46.86 | 90.92 |
| 2009 | 6.846 | 13.176 | 71.32 | 109.00 |
| 2010 | 5.023 | 12.046 | 68.24 | 104.67 |

 Table A5: Survey data for Gulf of Maine winter flounder (Nitschke, 2011).

Table A6a: NEFSC spring survey catch-at-age data for Gulf of Maine winter flounder (Nitschke, 2011).

| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8+ |
|------|---------|---------|---------|---------|---------|--------|--------|--------|
| 1982 | 92.06 | 1075.75 | 1900.83 | 474.97 | 570.39 | 62.23 | 0.00 | 116.13 |
| 1983 | 229.12 | 401.15 | 2462.32 | 1546.13 | 918.71 | 560.03 | 654.61 | 149.20 |
| 1984 | 117.19 | 640.90 | 901.25 | 554.72 | 315.92 | 92.45 | 154.91 | 107.51 |
| 1985 | 3.36 | 289.22 | 823.35 | 330.86 | 329.13 | 49.86 | 86.58 | 28.77 |
| 1986 | 17.96 | 433.05 | 217.59 | 308.31 | 54.06 | 202.14 | 59.71 | 18.13 |
| 1987 | 81.71 | 891.46 | 1480.03 | 368.52 | 187.09 | 32.68 | 66.93 | 30.61 |
| 1988 | 332.32 | 610.85 | 1895.85 | 706.61 | 190.39 | 82.21 | 29.61 | 12.03 |
| 1989 | 0.00 | 260.85 | 636.15 | 586.17 | 366.68 | 64.58 | 96.26 | 69.40 |
| 1990 | 12.82 | 448.05 | 1042.22 | 522.76 | 487.56 | 235.44 | 4.20 | 277.58 |
| 1991 | 34.70 | 619.24 | 985.48 | 540.22 | 285.31 | 54.34 | 8.62 | 0.00 |
| 1992 | 153.40 | 577.22 | 533.12 | 529.81 | 270.53 | 96.15 | 34.81 | 5.71 |
| 1993 | 0.00 | 250.89 | 345.92 | 148.98 | 98.55 | 9.51 | 17.18 | 0.00 |
| 1994 | 13.49 | 403.22 | 645.77 | 470.88 | 310.94 | 103.70 | 0.00 | 0.00 |
| 1995 | 161.96 | 1226.23 | 3090.63 | 1658.95 | 493.72 | 49.30 | 138.51 | 0.00 |
| 1996 | 39.12 | 180.65 | 538.43 | 509.44 | 240.20 | 14.83 | 8.28 | 0.00 |
| 1997 | 28.93 | 284.63 | 413.07 | 499.20 | 249.38 | 59.71 | 18.08 | 17.18 |
| 1998 | 58.31 | 328.96 | 335.67 | 269.41 | 118.20 | 5.32 | 0.00 | 3.97 |
| 1999 | 172.59 | 654.05 | 1276.04 | 940.03 | 398.47 | 183.79 | 18.13 | 0.00 |
| 2000 | 85.68 | 859.33 | 2136.77 | 1399.95 | 900.91 | 330.30 | 65.87 | 32.12 |
| 2001 | 39.40 | 289.84 | 787.19 | 833.64 | 462.88 | 333.04 | 121.50 | 66.15 |
| 2002 | 89.04 | 914.29 | 1670.48 | 1999.27 | 1280.52 | 513.98 | 188.71 | 96.54 |
| 2003 | 65.42 | 356.38 | 1203.79 | 1294.40 | 895.20 | 430.20 | 77.06 | 64.47 |
| 2004 | 299.30 | 466.35 | 494.33 | 414.36 | 186.42 | 209.70 | 100.51 | 0.00 |
| 2005 | 64.08 | 866.55 | 1278.73 | 789.99 | 438.54 | 288.94 | 102.41 | 43.15 |
| 2006 | 35.37 | 126.48 | 1065.67 | 664.02 | 332.99 | 85.01 | 25.86 | 0.00 |
| 2007 | 70.18 | 287.04 | 349.44 | 418.44 | 217.81 | 38.73 | 17.52 | 0.00 |
| 2008 | 1524.69 | 2335.33 | 1503.76 | 654.45 | 358.68 | 73.93 | 9.29 | 0.00 |
| 2009 | 33.63 | 618.24 | 1489.88 | 1100.43 | 474.02 | 69.00 | 41.86 | 4.20 |
| 2010 | 20.32 | 158.60 | 819.32 | 752.16 | 685.34 | 316.42 | 39.51 | 19.59 |

| | - | | | | | | | |
|------|---------|---------|---------|---------|---------|--------|--------|-------|
| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8+ |
| 1982 | 166.83 | 636.37 | 971.76 | 230.63 | 117.64 | 46.56 | 153.90 | 27.09 |
| 1983 | 1198.31 | 2012.87 | 1743.29 | 564.01 | 151.83 | 59.60 | 36.66 | 0.00 |
| 1984 | 250.50 | 1310.80 | 935.83 | 1216.16 | 332.60 | 124.30 | 61.90 | 95.31 |
| 1985 | 728.04 | 1533.42 | 1075.86 | 641.74 | 182.78 | 52.33 | 60.50 | 0.00 |
| 1986 | 16.85 | 403.67 | 645.88 | 272.60 | 30.61 | 11.42 | 0.00 | 18.92 |
| 1987 | 43.43 | 255.37 | 474.91 | 106.11 | 10.63 | 0.00 | 0.00 | 7.84 |
| 1988 | 237.79 | 572.96 | 338.53 | 394.66 | 85.91 | 30.89 | 18.13 | 0.00 |
| 1989 | 259.11 | 2015.33 | 792.01 | 419.62 | 52.66 | 37.72 | 0.00 | 6.27 |
| 1990 | 53.22 | 1039.03 | 610.79 | 221.90 | 30.61 | 12.03 | 6.04 | 0.00 |
| 1991 | 1452.33 | 1585.02 | 607.55 | 215.52 | 17.01 | 26.19 | 16.68 | 0.00 |
| 1992 | 1073.90 | 2072.97 | 1341.52 | 913.06 | 424.66 | 8.28 | 12.09 | 0.00 |
| 1993 | 927.61 | 1765.90 | 1015.75 | 385.09 | 130.45 | 5.65 | 0.00 | 0.00 |
| 1994 | 208.97 | 1288.30 | 846.18 | 354.03 | 22.05 | 5.65 | 0.00 | 0.00 |
| 1995 | 200.97 | 865.54 | 869.63 | 563.11 | 81.60 | 86.02 | 0.00 | 0.00 |
| 1996 | 987.88 | 1328.70 | 1440.52 | 1472.48 | 334.78 | 80.81 | 0.00 | 0.00 |
| 1997 | 231.19 | 2418.72 | 1787.72 | 823.63 | 320.68 | 18.80 | 0.00 | 0.00 |
| 1998 | 124.41 | 498.25 | 630.83 | 436.13 | 77.96 | 33.24 | 0.00 | 0.00 |
| 1999 | 453.37 | 1552.06 | 2040.57 | 1595.32 | 381.06 | 81.32 | 8.00 | 0.00 |
| 2000 | 349.16 | 1134.00 | 2238.63 | 1980.58 | 780.70 | 535.30 | 91.73 | 0.00 |
| 2001 | 200.58 | 927.38 | 1451.49 | 1564.59 | 539.55 | 203.93 | 23.73 | 5.93 |
| 2002 | 374.90 | 1535.49 | 1921.20 | 1317.96 | 698.88 | 109.52 | 11.70 | 13.32 |
| 2003 | 310.55 | 1779.16 | 1912.69 | 1004.00 | 562.33 | 111.15 | 18.24 | 0.00 |
| 2004 | 162.91 | 510.73 | 596.58 | 107.28 | 93.68 | 16.68 | 36.82 | 21.71 |
| 2005 | 699.89 | 1714.19 | 1313.43 | 751.88 | 327.61 | 54.51 | 30.67 | 36.49 |
| 2006 | 361.92 | 589.64 | 1718.72 | 758.82 | 490.53 | 22.22 | 41.25 | 0.00 |
| 2007 | 434.28 | 1174.69 | 760.78 | 774.43 | 315.69 | 109.30 | 0.00 | 0.00 |
| 2008 | 257.83 | 1391.66 | 2267.90 | 1873.80 | 1145.37 | 485.99 | 0.00 | 31.56 |
| 2009 | 80.31 | 1558.66 | 2246.74 | 1757.39 | 1320.20 | 382.96 | 20.99 | 6.16 |
| 2010 | 21.77 | 576.66 | 1908.49 | 2241.48 | 1448.19 | 307.47 | 190.11 | 46.84 |

Table A6b: NEFSC fall survey catch-at-age data for Gulf of Maine winter flounder (Nitschke,2011).

Table A6c: Massachusetts spring survey catch-at-age data for Gulf of Maine winter flounder(Nitschke, 2011).

| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8+ |
|------|---------|----------|---------|---------|---------|---------|--------|--------|
| 1982 | 1658.16 | 6361.20 | 1836.79 | 1947.30 | 419.94 | 111.03 | 354.05 | 14.45 |
| 1983 | 3175.87 | 6278.94 | 6590.91 | 4132.55 | 1984.33 | 537.40 | 211.29 | 270.09 |
| 1984 | 1309.41 | 5596.43 | 3427.98 | 2620.57 | 907.69 | 55.09 | 216.20 | 82.31 |
| 1985 | 2136.45 | 1672.36 | 3405.67 | 2706.69 | 1008.58 | 146.04 | 49.75 | 23.13 |
| 1986 | 2295.95 | 3780.69 | 5293.35 | 2403.82 | 349.06 | 43.77 | 16.58 | 39.63 |
| 1987 | 3593.98 | 3635.32 | 7003.94 | 2556.53 | 169.83 | 334.75 | 85.54 | 182.30 |
| 1988 | 1650.94 | 3169.97 | 3910.95 | 2118.61 | 170.88 | 56.20 | 13.38 | 45.67 |
| 1989 | 2065.50 | 4331.80 | 3825.33 | 2021.71 | 823.53 | 166.82 | 28.18 | 75.85 |
| 1990 | 3265.41 | 4208.90 | 6244.24 | 2282.77 | 471.16 | 312.09 | 108.94 | 38.78 |
| 1991 | 984.48 | 3502.56 | 2550.99 | 1649.27 | 683.00 | 110.68 | 53.05 | 103.32 |
| 1992 | 4447.96 | 6709.79 | 2828.85 | 1562.76 | 476.95 | 173.95 | 39.07 | 48.72 |
| 1993 | 4039.88 | 7104.62 | 2796.09 | 1358.03 | 435.10 | 264.34 | 49.42 | 38.14 |
| 1994 | 3310.73 | 7781.04 | 3075.77 | 1000.02 | 169.00 | 26.08 | 36.60 | 21.32 |
| 1995 | 4474.23 | 7433.04 | 4751.29 | 1288.26 | 294.75 | 117.69 | 32.78 | 32.10 |
| 1996 | 3212.26 | 5900.09 | 3246.03 | 1531.28 | 419.30 | 165.27 | 42.62 | 17.48 |
| 1997 | 3199.00 | 7320.28 | 3758.75 | 1838.01 | 1030.33 | 220.03 | 134.86 | 105.44 |
| 1998 | 2106.16 | 5871.00 | 4368.13 | 2399.67 | 833.70 | 248.73 | 169.48 | 37.69 |
| 1999 | 3181.83 | 5455.98 | 4427.92 | 2024.19 | 1045.28 | 268.24 | 134.22 | 116.41 |
| 2000 | 5997.82 | 13694.80 | 6182.92 | 3527.20 | 2279.87 | 1248.05 | 381.06 | 128.22 |
| 2001 | 3038.06 | 2156.92 | 5664.30 | 4172.67 | 1605.77 | 1067.98 | 528.95 | 268.38 |
| 2002 | 1891.06 | 6962.83 | 4197.19 | 3884.95 | 1482.33 | 263.39 | 94.94 | 2.54 |
| 2003 | 3172.08 | 6338.79 | 3738.01 | 2264.07 | 1262.44 | 353.59 | 108.26 | 18.25 |
| 2004 | 5569.03 | 6461.18 | 1671.73 | 1208.82 | 911.14 | 381.53 | 70.33 | 37.83 |
| 2005 | 7223.85 | 8227.77 | 2691.42 | 870.50 | 305.58 | 57.54 | 7.07 | 5.98 |
| 2006 | 4302.98 | 8758.47 | 2948.09 | 1189.54 | 331.10 | 70.95 | 26.99 | 10.00 |
| 2007 | 2302.69 | 4893.18 | 2081.50 | 1254.46 | 398.77 | 94.72 | 13.44 | 8.78 |
| 2008 | 2072.08 | 4453.26 | 1452.02 | 1133.50 | 417.03 | 93.32 | 27.65 | 13.15 |
| 2009 | 2115.48 | 4797.99 | 3989.67 | 1995.28 | 1290.75 | 364.95 | 103.95 | 45.75 |
| 2010 | 1832.75 | 3890.83 | 3509.46 | 2881.02 | 1191.91 | 539.98 | 194.14 | 28.37 |

Table A6d: Massachusetts fall survey catch-at-age data for Gulf of Maine winter flounder(Nitschke, 2011).

| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8+ |
|------|----------|---------|---------|---------|---------|--------|--------|-------|
| 1982 | 9419.66 | 7334.77 | 4407.41 | 810.44 | 147.11 | 46.47 | 20.97 | 20.49 |
| 1983 | 8909.33 | 3589.56 | 2474.79 | 572.97 | 229.08 | 14.04 | 1.57 | 13.73 |
| 1984 | 1715.39 | 2715.77 | 1434.21 | 1640.61 | 449.45 | 121.07 | 53.19 | 8.84 |
| 1985 | 4897.43 | 2810.59 | 1411.26 | 638.41 | 160.34 | 38.18 | 18.47 | 8.58 |
| 1986 | 3738.84 | 3230.42 | 1830.09 | 319.49 | 43.50 | 0.00 | 0.00 | 21.28 |
| 1987 | 3325.39 | 4315.82 | 3177.09 | 249.97 | 9.26 | 15.26 | 23.15 | 24.37 |
| 1988 | 2789.74 | 3194.71 | 935.84 | 672.78 | 185.46 | 99.42 | 34.14 | 0.00 |
| 1989 | 2794.61 | 3609.79 | 1286.50 | 292.09 | 65.44 | 22.56 | 10.62 | 10.62 |
| 1990 | 1801.47 | 9234.03 | 2325.97 | 532.45 | 48.99 | 0.00 | 0.00 | 7.15 |
| 1991 | 10419.18 | 6327.18 | 2900.09 | 604.12 | 8.99 | 70.76 | 26.93 | 0.00 |
| 1992 | 9367.51 | 4532.62 | 1891.98 | 1295.20 | 675.75 | 67.61 | 21.44 | 57.21 |
| 1993 | 7523.20 | 7769.60 | 2747.19 | 747.78 | 331.78 | 65.28 | 21.44 | 21.46 |
| 1994 | 2918.62 | 6752.77 | 3179.56 | 1042.23 | 47.30 | 0.00 | 5.38 | 5.46 |
| 1995 | 5419.59 | 4880.19 | 3341.76 | 1844.44 | 133.38 | 76.55 | 10.93 | 34.33 |
| 1996 | 7524.31 | 3352.89 | 2575.63 | 1884.97 | 265.84 | 92.20 | 4.78 | 4.78 |
| 1997 | 4814.83 | 6418.38 | 3467.90 | 1051.98 | 317.18 | 14.93 | 0.00 | 0.00 |
| 1998 | 8603.17 | 5826.52 | 3839.39 | 1490.50 | 272.57 | 155.48 | 15.22 | 20.72 |
| 1999 | 7886.42 | 8744.32 | 4914.05 | 3132.82 | 783.29 | 126.35 | 15.71 | 26.70 |
| 2000 | 5374.73 | 5949.39 | 4929.16 | 2799.49 | 787.06 | 559.15 | 132.26 | 0.00 |
| 2001 | 6126.97 | 3548.97 | 2918.46 | 2868.44 | 787.16 | 327.31 | 37.26 | 22.14 |
| 2002 | 3776.65 | 4675.99 | 2613.62 | 1531.07 | 686.63 | 93.81 | 16.02 | 15.98 |
| 2003 | 10176.70 | 4439.24 | 2015.05 | 979.79 | 458.87 | 61.36 | 25.65 | 0.00 |
| 2004 | 11968.46 | 4887.41 | 3668.01 | 544.31 | 411.16 | 145.38 | 127.40 | 16.35 |
| 2005 | 5186.41 | 7090.88 | 2258.30 | 1090.90 | 435.57 | 40.02 | 31.92 | 38.62 |
| 2006 | 6248.84 | 4626.76 | 4821.72 | 1472.35 | 616.90 | 11.46 | 39.26 | 9.48 |
| 2007 | 7590.02 | 4281.27 | 1958.21 | 1358.60 | 422.81 | 107.75 | 14.14 | 2.76 |
| 2008 | 5706.92 | 5761.15 | 3592.51 | 2148.16 | 1096.47 | 307.90 | 0.00 | 35.38 |
| 2009 | 4210.96 | 9523.65 | 4708.05 | 2278.75 | 1288.61 | 365.94 | 16.37 | 34.91 |
| 2010 | 4923.51 | 6220.98 | 4294.42 | 3028.39 | 1596.86 | 618.12 | 341.24 | 66.76 |

Appendix B - The Age-Structured Production Model

The model used for these assessments is an Age-Structured Production Model (ASPM) (e.g. Hilborn, 1990). Models of this type fall within the more general class of Statistical Catch-at-Age Analyses. The approach used in an ASPM assessment involves constructing an age-structured model of the population dynamics and fitting it to the available abundance indices by maximising the likelihood function. The model equations and the general specifications of the model are described below, followed by details of the contributions to the (penalised) log-likelihood function from the different sources of data available and assumptions concerning the stock-recruitment relationship. Quasi-Newton minimization is used to minimize the total negative log-likelihood function (the package AD Model Builder[™], Otter Research, Ltd is used for this purpose).

B.1. Population dynamics

B.1.1 Numbers-at-age

The resource dynamics are modelled by the following set of population dynamics equations:

$$N_{y+1,1} = R_{y+1}$$
(B1)

$$N_{y+1,a+1} = \left(N_{y,a} \ e^{-M_{y,a}/2} - C_{y,a}\right) e^{-M_{y,a}/2} \qquad \text{for } 1 \le a \le m-2 \tag{B2}$$

$$N_{y+1,m} = \left(N_{y,m-1} e^{-M_{y,m-1}/2} - C_{y,m-1}\right) e^{-M_{y,m-1}/2} + \left(N_{y,m} e^{-M_{y,m}/2} - C_{y,m}\right) e^{-M_{y,m}/2}$$
(B3)

where

$$N_{y,a}$$
 is the number of fish of age *a* at the start of year *y* (which refers to a calendar year),

 R_y is the recruitment (number of 1-year-old fish) at the start of year y,

 $M_{y,a}$ denotes the natural mortality rate for fish of age *a* in year *y*,

 $C_{y,a}$ is the predicted number of fish of age *a* caught in year *y*, and

m is the maximum age considered (age 8 here) (taken to be a plus-group).

B.1.2. Recruitment

The number of recruits at the start of year *y* is assumed to follow a Beverton-Holt stock-recruit curve, and allowing for annual fluctuation about the deterministic relationship:

$$R_{y} = \frac{\alpha B_{y}^{sp}}{\beta + B_{y}^{sp}} e^{(\varsigma_{y} - (\sigma_{R}^{\perp})^{2}/2)}$$
(B4)

where

 α and β are spawning biomass-recruitment relationship parameters,

- ς_y reflects fluctuation about the expected recruitment for year y, which is assumed to be normally distributed with standard deviation $\sigma_R = 0.5$
- B_{v}^{sp} is the spawning biomass, computed as:

$$B_{y}^{sp} = \sum_{a=1}^{m} f_{y,a} w_{y,a}^{strt} N_{y,a} e^{-M_{y,a}\delta}$$
(B5)

where

 $w_{v,a}^{sp}$ is the mass of fish of age a during spawning, and

 $f_{y,a}$ is the proportion of fish of age *a* that are mature,

 δ is the proportion of the natural mortality that occurs before spawning (0.25 here).

B.1.3. Total catch and catches-at-age

The catch by mass in year y is given by:

$$C_{y} = \sum_{a=1}^{m} w_{y,a}^{mid} C_{y,a} = \sum_{a=1}^{m} w_{y,a}^{mid} N_{y,a} e^{-M_{y,a}/2} S_{y,a} F_{y}$$
(B6)

where

- $w_{y,a}^{mid}$ denotes the mass of fish of age *a* landed in year *y*,
- $C_{y,a}$ is the catch-at-age, i.e. the number of fish of age *a*, caught in year *y*,
- $S_{y,a}$ is the commercial selectivity (i.e. combination of availability and vulnerability to fishing gear) at age *a* for year *y*; when $S_{y,a} = 1$, the age-class *a* is said to be fully selected, and
- F_{v} is the proportion of a fully selected age class that is fished.

The model estimate of the mid-year exploitable ("available") component of biomass is calculated by converting the numbers-at-age into mid-year mass-at-age (using the individual weights of the landed fish) and applying natural and fishing mortality for half the year:

$$B_{y}^{ex} = \sum_{a=1}^{m} w_{y,a}^{mid} S_{y,a} N_{y,a} e^{-M_{y,a}/2} (1 - S_{y,a} F_{y}/2)$$
(B7)

For survey estimates (in numbers):

$$N_{y}^{surv,i} = \sum_{a=1}^{m} S_{a}^{i} N_{y,a} e^{-M_{y,a} \frac{\overline{\sigma}^{i}}{12}} \left(1 - S_{y,a} F_{y} \frac{\overline{\sigma}^{1}}{12} \right)$$
(B8)

where

 S_a^i is the survey selectivity for age *a* and survey *i*,

 $\overline{\omega}^i$ is the month in which survey *i* has taken place.

B.1.4. Initial conditions

For the first year (y_0) considered in the model therefore, the stock is assumed to be at a level $B_{y_0}^{sp}$ (estimated in the model fitting procedure), with the starting age structure:

$$N_{y_0,a} = R_{start} N_{start,a} \qquad \qquad \text{for } 1 \le a \le m \tag{B9}$$

where

$$N_{start,1} = 1 \tag{B10}$$

$$N_{start,a} = N_{start,a-1} e^{-M_{y_0,a-1}} (1 - \phi S_{y_0,a-1}) \qquad \text{for } 2 \le a \le m-1$$
(B11)

$$N_{start,m} = N_{start,m-1} e^{-M_{y_0,m-1}} (1 - \phi S_{y_0,m-1}) / (1 - e^{-M_{y_0,m}} (1 - \phi S_{y_0,m}))$$
(B12)

where ϕ characterises the average fishing proportion over the years immediately preceding y_0 .

B.2. The (penalised) likelihood function

The model is fit to survey abundance indices, and commercial and survey catch-at-age data to estimate model parameters (which may include residuals about the stock-recruitment function, the fishing selectivities, the annual catches or natural mortality, facilitated through the incorporation of the penalty functions described below). Contributions by each of these to the negative of the (penalised) log-likelihood ($-\ell nL$) are as follows.

B.2.1. Survey abundance data

The likelihood is calculated assuming that an observed survey index is log-normally distributed about its expected value:

$$I_{y}^{i} = \hat{I}_{y}^{i} \exp\left(\varepsilon_{y}^{i}\right) \quad \text{or} \quad \varepsilon_{y}^{i} = \ell n \left(I_{y}^{i}\right) - \ell n \left(\hat{I}_{y}^{i}\right) \tag{B13}$$

where

 I_{y}^{i} is the survey index for year y and series i,

 $\hat{I}_{y}^{i} = \hat{q}^{i} N_{y}^{surv,i}$ is the corresponding model estimate, where $N_{y}^{surv,i}$ is the model estimate, given by equation (B8),

$$\hat{q}^i$$
 is the constant of proportionality (catchability) for index *i*, and

$$\varepsilon_y^i$$
 from $N\left(0, \left(\sigma_y^i\right)^2\right)$.

For these analyses, selectivities are estimated as detailed in section B.3.1 below.

The contribution of the survey abundance data to the negative of the log-likelihood function (after removal of constants) is then given by:

$$-\ln L^{survey} = \sum_{i} \sum_{y} \left[\ln \left(\sigma_{y}^{i} \right) + \left(\varepsilon_{y}^{i} \right)^{2} / 2 \left(\sigma_{y}^{i} \right)^{2} \right]$$
(B14)

where

 σ_{y}^{i} is the standard deviation of the residuals for the logarithm of index *i* in year *y*.

Homoscedasticity of residuals is assumed, so that $\sigma_y^i = \sigma^i$ is estimated in the fitting procedure by its maximum likelihood value:

$$\hat{\sigma}^{i} = \sqrt{1/n_{i} \sum_{y} \left(\ell n(I_{y}^{i}) - \ell n(q^{i} N_{y}^{surv,i}) \right)^{2}}$$
(B15)

where

 n_i is the number of data points for survey index *i*.

The catchability coefficient q^i for survey index *i* is estimated by its maximum likelihood value:

$$\ell n \, \hat{q}^{i} = 1/n_{i} \sum_{y} \left(\ln I_{y}^{i} - \ln N_{y}^{surv,i} \right) \tag{B16}$$

To allow for first order serial correlation between the survey residuals, a serial correlation coefficient ρ^i would be estimated for each survey index:

$$\boldsymbol{\varepsilon}_{y}^{i} = \boldsymbol{\lambda}_{y}^{i} - \boldsymbol{\rho}\boldsymbol{\lambda}_{y-1}^{i} \tag{B17}$$

where

$$\lambda_{y}^{i} = \ell n \left(I_{y}^{i} \right) - \ell n \left(\hat{I}_{y}^{i} \right)$$

and the summation in equation (B.16) extends over one less year.

B.2.2. Commercial catches-at-age

The contribution of the catch-at-age data to the negative of the log-likelihood function under the assumption of an "adjusted" lognormal error distribution is given by:

$$- \ln L^{CAA} = w^{CAA} \sum_{y} \sum_{a} \left[\ln \left(\sigma_{com} / \sqrt{p_{y,a}} \right) + p_{y,a} \left(\ln p_{y,a} - \ln \hat{p}_{y,a} \right)^2 / 2 \left(\sigma_{com} \right)^2 \right]$$
(B18)

where

 w^{comCAA} is a multiplicative factor to downweight the commercial CAA likelihood,

 $p_{y,a} = C_{y,a} / \sum_{a'} C_{y,a'}$ is the observed proportion of fish caught in year y that are of age a,

 $\hat{p}_{y,a} = \hat{C}_{y,a} / \sum_{a'} \hat{C}_{y,a'}$ is the model-predicted proportion of fish caught in year y that are of age *a*, where

$$\hat{C}_{y,a} = N_{y,a} \ e^{-M_{y,a}/2} \ S_{y,a} \ F_y \tag{B19}$$

and

 σ_{com} is the standard deviation associated with the catch-at-age data, which is estimated in the fitting procedure by:

$$\hat{\sigma}_{com} = \sqrt{\sum_{y} \sum_{a} p_{y,a} \left(\ln p_{y,a} - \ln \hat{p}_{y,a} \right)^2 / \sum_{y} \sum_{a} 1}$$
(B20)

B.2.3. Survey catches-at-age

The survey catches-at-age are incorporated into the negative of the log-likelihood in an analogous manner to the commercial catches-at-age (thus they are also weighted by a factor $w^{survCAA}$), assuming an adjusted log-normal error distribution (equation (B18)) where:

$$p_{y,a} = C_{y,a}^{surv} / \sum_{a'} C_{y,a'}^{surv}$$
 is the observed proportion of fish of age *a* in year *y*, with

$$C_{y,a}^{surv,i} = S_{a}^{i} N_{y,a} e^{-M_{y,a} \frac{\overline{\omega}^{i}}{12}} \left(1 - S_{y,a} F_{y} \frac{\overline{\omega}^{1}}{12} \right)$$
(B21)

 $\hat{p}_{y,a}$ is the expected proportion of fish of age *a* in year *y* in the survey.

B.2.4. Stock-recruitment function residuals

The stock-recruitment residuals are assumed to be log-normally distributed. Thus, the contribution of the recruitment residuals to the negative of the (now penalised) log-likelihood function is given by:

$$- \ln L^{SRpen} = \sum_{y=y1}^{1988} \left[\frac{\varepsilon_y^2}{2(\sigma_R^1)^2} \right] + \sum_{1989}^{y^2} \left[\frac{\varepsilon_y^2}{2(\sigma_R^2)^2} \right]$$
(B22)

where

$$\varepsilon_{y}$$
 from $N(0, (\sigma_{R})^{2})$

B.3. Model parameters

B.3.1. Fishing selectivity-at-age:

The commercial selectivity is estimated separately for ages 1 to 4 and is assumed to be flat for ages 5 and above (except for case 8) for which selectivity is also estimated separately for ages 5 and above. The survey fishing selectivities are estimated separately for ages 1 to a_{plus} (the plus group age) and flat thereafter. a_{plus} =7 for the spring surveys and 6 for the fall surveys. B.3.2.: Other parameters reported in Table 1 and elsewhere

$$\sigma_{R}_{-} \text{out} = \sum_{y=y_1}^{y_2} (\varsigma_y)^2 / \sum_{y=y_1}^{y_2} 1$$
(B23)

where *y*1=1982 and *y*2=2010.

Calculation of MSY

ſ

The equilibrium catch for a fully selected fishing proportion *F* is calculated as:

$$C(F) = \sum_{a} w_{a}^{mid} S_{a} F N_{a} (F) e^{-(M_{a}/2)}$$
(B24)

where
$$w_a^{mid} = \sum_{y_1=2006}^{2010} w_{y,a}^{mid} / 5$$
, $S_a = S_{2010,a}$ and $M_a = M_{2010,a}$

and where numbers-at-age *a* are given by:

$$N_{a}(F) = \begin{cases} R_{1}(F) & \text{for } a = 1\\ N_{a-1}(F)e^{-M_{a-1}}(1 - S_{a-1}F) & \text{for } 1 < a < m\\ \frac{N_{m-1}(F)e^{-M_{m-1}}(1 - S_{m-1}F)}{(1 - e^{-M_{m}}(1 - S_{m}F))} & \text{for } a = m \end{cases}$$
(B25)

where

$$R_1(F) = \frac{\alpha B^{sp}(F)}{\beta + B^{sp}(F)}$$
(B26)

The maximum of C(F) is then found by searching over F to give F_{MSY} , with the associated spawning biomass and yield given by

$$B_{\rm MSY}^{sp} = \sum_{a} f_a w_a^{strt} N_a (F_{\rm MSY}) e^{-M_a \delta}$$
(B27)

$$MSY = \sum_{a} w_a^{mid} S_a F_{\rm MSY} N_a \left(F_{\rm MSY} \right) e^{-(M_a/2)}$$
(B28)

where
$$w_a^{strt} = \sum_{y_1=2006}^{2010} w_{y,a}^{strt} / 5$$
 and $f_a = \sum_{y_1=2006}^{2010} f_{y,a} / 5$

ADDITIONAL REFERENCE

Hilborn, R. 1990. Estimating the parameters of full age-structured models from catch and abundance data. International North Pacific Fisheries Commission Bulletin, 50: 207-213.